Announcements and credibility under inflation targeting

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Abstract

We inspect how inflation target announcements are instrumental in building central bank credibility and shaping inflation expectations. Investigating the role of announcements by using a time varying credibility measure, we find that both the accuracy and the frequency of inflation announcements have a positive impact on how much attention the public pays to target announcements.

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1. Introduction

Adoption of inflation targeting (IT) is aimed at moderating inflation expectations by not only providing discipline in the setting of monetary policy, but also by improving the communication between the policy makers and the public. If the central bank does not gain the necessary trust to hone the public’s inflation expectations toward its announced target, then expectations respond only slowly to the target rate, and the monetary authority fails to eliminate the inflation bias of economic agents. Bernanke, in his FRB speech, emphasizes the importance of central bank credibility “...Clearly there are limits to what talk can achieve; ultimately, talk must be backed up by action, in the form of successful policies... Credibility is not a permanent characteristic of a central bank; it must be continuously earned...”\textsuperscript{1} In our study, we investigate how the announcements shape inflation expectations and how the IT central bank’s performance in hitting its announced target influences its credibility.

The literature on credibility of announcements (Cukierman, 1992; Faust and Svensson, 2001; Walsh, 1999, 2003) argues that there are two gains from making announcements, 1) the central bank is able to respond to the supply shocks (without distorting the public’s inflation expectation) that only itself observes, and more importantly 2) the public can improve its forecast of the central bank’s true preferences. In environments where the public is uncertain about the (weights of the) central bank’s reaction function, they will use the target announcements as additional information in forming their inflation expectations. The empirical studies about IT’s impact on inflation expectations find that credibility and control over inflation expectations cannot be achieved immediately. Johnson (2002), using a panel of 11 countries, shows that IT leads to declines in inflation forecasts without affecting their dispersion while Fraga et al. (2003) argue

that IT central banks in emerging markets (with imperfect credibility) have to work harder toward this goal. Only forceful reactions to economic shocks by the central bank will force agents’ inflation expectations not to significantly deviate from the target and increase the credibility of the central bank.

In our analysis, we inspect the impact of target announcements to see how they help the public to learn about the central bank’s objectives. The hypothesis we would like to test is how “credible” and “stable” announcements change the public’s learning behavior. As the central bank delivers on its announced targets or relays as much information as possible to the public, the public should concentrate their inflation expectations around the announced target. In testing our claims, we use survey data from two early adopters of IT, United Kingdom (UK) and New Zealand (NZ), with sufficient length of data under IT and significant differences between the frequencies of their inflation target announcements. Both countries adopted the IT regime in early 1990s (1993 for UK and 1990 for NZ), giving us at least 13 years of data for our analysis. Aside from the sample length, UK is very different than NZ in having a very stable inflation target with only one change in their target while NZ changed its targets quite frequently (40 changes in the sample period).

Using inflation forecasts from both countries and state space methodology, we find that forecasters give more weight to the announcements when they are more accurate and frequently changing. In the next section, we elaborate on the model while Section 3 displays the estimation results. The last section concludes.

2. Model and methodology

In our model, we follow Walsh (1999) and assume that in an environment with uncertain policy objectives, central banks realize the potential of their announcements to influence inflation expectations of private agents. These announcements reveal more information to the public on the central bank’s intentions and also discipline the central bank due to its continuous updating (of $\theta_t$) during the process of expectations formation. Using a more general function, $\theta_t = -f(\pi_{t-1} - \pi^e_{t-1})$, approximating it in the second-order around zero ($E(\theta_t) = -f(0) - E(\pi_{t-1} - \pi^e_{t-1})f' - E(\pi_{t-1} - \pi^e_{t-1})^2f''/2$) and calculating the variance term using a moving average window enable us to also include previous performances of the central bank in the reputation function. Specifically, we expect to find increasing $\theta_t$ as the central bank gains more credibility by hitting its target repeatedly, so the expected signs for $f'$ and $f''$ are negative and positive, respectively, while $-f(0)$ should be 1.

In our estimations we use a combination of survey and actual inflation data for UK and NZ. The expectations data for retail price inflation in UK and CPI inflation in NZ are obtained from Consensus Economics. Consensus forecasts are monthly reported expectations for the end of current year and the next year. Using them and realized inflation series, we derive monthly expected rate of inflation for the next 12-months as in Johnson (2003). The calculation results are displayed in Figs. 1 and 2, which show a tight fit in UK of expectations to actual inflation especially after the switch to IT whereas in NZ, the situation is not that clear-cut.

Since the expected inflation and announcement data are observed, we use state space estimation, specifically the Kalman filter method, to estimate our time varying credibility parameter. Kalman filter is quite appropriate for our purposes due to its continuous updating (of $\theta_t$) during the process of expectations formation. We choose the initial parameter values as 1 for $\theta_0$, and 0.9 for its persistence; and finally we determine the error variances, $\sigma^2_\varepsilon$ and $\sigma^2_\xi$ via a grid search and pick the values that give us the minimum regression standard error.

3. Results

As mentioned earlier, an increase in credibility via good inflation performance should lead to stable inflation expectations.
and tighter concentration around the announcements, hence an increase in $\theta_t$. Using the estimated time-varying values for $\theta_t$, we calculate the predicted inflation series, displayed in Figs. 3 and 4. An interesting result of this initial step is the deviation of model predictions from UK inflation survey values around 1997. This occurs right around the Bank of England’s (BoE) independence, pointing to increased public uncertainty about the new objective function of the central bank.

Next, we analyze how and why the public changes their learning parameter in response to the central bank’s accuracy of...

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Fig. 1. Inflation and expectations in UK.

Fig. 2. Inflation and expectations in New Zealand.
announcements in mimicking actual inflation and their variability. Therefore, the variables used in the regression are the terms in the approximated credibility function, namely difference of inflation from the announcement $|\pi_{t-1}^{\text{UK}} - \pi_{t-1}^{\text{NZ}}|$ and the variation of this differential calculated from a 1-year moving window, $\text{var}(\pi^{\text{UK}} - \pi)$. We estimate this system of two countries with the help of a dummy variable $D_{\text{UK}}$, which takes the value of 1 for UK and 0 for NZ. We run a second estimation with a decomposed version of the
The variance term, into individual variances and covariance. The estimation results are displayed in the second and third columns of Table 1. The results show that (i) above 35% of variation in theta is explained by the two moments of the inflation-announcement differential, and (ii) credibility increases with more accurate announcements and less variation in this accuracy. The significantly different first moment coefficients show that the agents in UK value the information content of the past record of BoE in hitting its targets more than the forecasters in NZ. Decomposition of the variance term illustrates that the more frequent the announcements are the more people pay attention to them. In other words, since BoE has had only one target rate change in the 13 years under IT regime, the past deviations of actual inflation from the target are frequently used by the public in updating their expectations. On the other hand, in NZ, since the targets change extremely frequently, the public focuses more on the announcements than the actual inflation itself, apparent from the positive coefficient on the var(πa − π) parameter. In short, our estimations show that the announcements carry a large information content in IT regime and their accuracy and frequency influence the extent the public believes in the central bank’s announcements and acts accordingly.

4. Conclusion

In our study, we examine the credibility performance of central banks under inflation targeting. We use inflation surveys from the UK and New Zealand due to significant dissimilarity between the two countries in the operational application of IT monetary policy. Estimating our model with state space representation of a time varying credibility measure shows that both the accuracy and the frequency of inflation announcements have a positive effect on how much attention the public pays to the target announcements.

References


<table>
<thead>
<tr>
<th>Table 1</th>
<th>Empirical analysis of the time varying parameter $\theta_t$</th>
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<tbody>
<tr>
<td>Coeff.</td>
<td>Coeff. (variance broken down)</td>
</tr>
<tr>
<td>Constant</td>
<td>0.12** (0.02)</td>
</tr>
<tr>
<td>$D_{UK}$</td>
<td>0.17** (0.03)</td>
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<td>$</td>
<td>\pi^a - \pi</td>
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<tr>
<td>$</td>
<td>\pi^a - \pi</td>
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<tr>
<td>var($\pi^a - \pi$)</td>
<td>0.03** (0.01)</td>
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<tr>
<td>var($\pi^a - \pi</td>
<td>D_{UK}$</td>
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<tr>
<td>var($\pi$)</td>
<td>0.129* (0.066)</td>
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<tr>
<td>var($\pi</td>
<td>D_{UK}$</td>
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<tr>
<td>var($\pi</td>
<td>D_{UK}$</td>
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<tr>
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<td>$R^2$</td>
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<tr>
<td>$F$</td>
<td>29.91</td>
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** (*) indicates 95% (90%) significance level. Since $D_{UK}$ is 1 for UK and 0 for NZ, NZ sets the benchmark while interactive terms represent UK. Due to the stability of the target in UK, the dummy and interactive variance term are very highly correlated. Separate regressions of UK result in even a higher positive coefficient for the variance term. Covariance term is left out of the regression due to its high (0.98) correlation with NZ target variance.

5 The covariance term is left out of the regression due to its high collinearity with the NZ target variance.
6 Addition of output gap brings another 5% explanation of the variation in theta.
7 The magnitude of the interactive term coefficient for announcement variance in UK is due to the almost nonexistent variation in the target.